STATISTICAL AND MATHEMATICAL FOUNDATIONS



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MODULE LEARNING OBJECTIVES

In this module, we provide, in very broad terms, some fundamental mathematical and statistical background required for data analysis and model building with practical applications.

Participants will become acquainted will key concepts, to allow for future learning.

This introduction is not meant to replace formal training and is at best **incomplete**; please consult the references for further details.





OUTLINE

- 1. Modeling
- 2. Distributions
- 3. Central Limit Theorem
- 4. Estimation
- 5. Bayes' Theorem
- 6. Matrix Algebra

- 7. Eigenvalues and Eigenvectors
- 8. Optimization
- 9. Sampling Methods
- 10. Confidence Intervals
- **11.** Hypothesis Testing
- 12. Regression Analysis



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MODELING

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LEARNING OBJECTIVES

Understand the difference between modeling from first principles and statistical modeling.

Working knowledge of the modeling process.

Increase awareness of modeling pitfalls and challenges.







Theory

Identification of details relevant to **description** and **translation** of realworld objects into model variables Model





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MODELS IN GENERAL

First principles modeling

- examine a system
- write down a set of rules/equations that describe the essence of the system
- ignore complicating details that are "less" important

Statistical modeling

- typically a set of equations with parameters
- parameters are learned (model is "trained") using multiple data observations
- data sample vs. population





MODELING HEURISTICS

In a sense, modeling is a **straightforward** (and **formulaic**?) process, guided by **intuition** and **experience** at each step.

Basic steps in building a statistical model:

defining the goals

- what are we trying to achieve?
- under what situations will the model be used and what is the outcome we are trying to predict?

gathering data

- what data is available?
- how many records will we have?
- generally, modelers want as much data as possible





MODELING HEURISTICS

Basic steps in building a statistical model: (continued)

deciding on the model structure

- should we run a linear regression, logistic regression, or a nonlinear model? Which kind?
- choices of model structure require experience and deep knowledge of the strength and weaknesses of each technique

preparing the data

- assemble data into appropriate form for the model
- encode the data into inputs, using expert knowledge as much as possible
- separate the data into the desired training, testing, and validation sets





MODELING HEURISTICS

Basic steps in building a statistical model: (continued)

selecting and removing features

- variables are examined for model importance and selected or eliminated
- a list of candidate appropriate variables are ordered by importance

building candidate models

- begin with baseline linear models and try to improve using more complex nonlinear models
- keep in mind the environment in which the model will be implemented

finalizing the model

- select among the candidates the most appropriate model to be implemented

implementing and monitoring

- embed the model into necessary system process; implement monitoring steps to examine the model performance





MODELING PITFALLS

Common pitfalls surrounding the modeling process:

defining the goals

- lack of clarity around problem definition
- lack of understanding of how and where the model will be used

gathering data

- using data that is too old or otherwise not relevant going forward
- not considering additional key data sources or data sets that might be available

deciding on the model structure

- using a modeling methodology that is not appropriate for the nature of the data (sizes, dimensions, noise...)





MODELING PITFALLS

Common pitfalls surrounding the modeling process: (continued)

preparing the data

- not cleaning or considering outliers
- not properly scaling data
- not giving enough thought to building special expert variables
- not having data from important categories of records

selecting and eliminating features

- keeping too many variables, making it hard for modeling, interpretation, implementation, or model maintenance
- too much reliance on simply eliminating correlated variables





MODELING PITFALLS

Common pitfalls surrounding the modeling process: (continued)

building candidate models

- overfitting
- not doing proper training/testing as one examines candidate models
- not doing a simpler linear regression to use as baseline

finalizing the model

- not rebuilding the final model optimally using all the appropriate data
- improperly selecting the final model without consideration to some implementation constraints

implementing and monitoring

- errors in implementation process: data input streams, variable encodings, algorithm mistakes
- not monitoring model performance





DISTRIBUTIONS

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LEARNING OBJECTIVES

What questions can you use to help you pick a model distribution for a data feature?

What are some commonly encountered pdfs?

What are the mean and variance of some common pdfs?

When do we need to use joint distributions?





DATA AND DISTRIBUTIONS

If a data feature can be characterized by a distribution, consider asking **four basic questions**:

- 1. Can the variable only take on **discrete** values? **continuous** values?
 - whether a taxpayer's file is audited or not is a *discrete* variable but the corrected amount from the audit is a *continuous* variable
- 2. Is the data distribution **symmetric**?
 - If not, in which **direction** does the asymmetry lie?
 - Are **right-** and **left-outliers** equally likely?

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DATA AND DISTRIBUTIONS

- 3. Does the variable have theoretical **upper** and **lower limits**?
 - Some items like age or height cannot be smaller than zero
 - Some items like operating margins cannot exceed a value (100% in this case)
- 4. How likely is it to observe **extreme values** in the distribution?
 - in some data, extreme values occur infrequently whereas in others, they occur more often

How would these questions have to change when dealing with joint distributions?





FUNDAMENTAL DISTRIBUTIONS

Empirical distributions are often approximated by **parametric distributions**, defined *via* a **probability density function** (pdf) and a set of parameters that must be learned from the data.

The basic distributions of data analysis are:

- the **uniform** distribution U(a,b) on the interval [a,b] or $U(x_1,...,x_n)$ on the discrete set $\{x_1,...,x_n\}$, potentially the simplest
- the **normal** distribution $N(\mu, \sigma^2)$ on the real line \mathbb{R} , possibly the most frequently used (not always aptly so)
- a wide variety of special distributions that are used in applications ranging from consumer modeling and finance to operation research (Poisson, exponential, log-normal, binomial, etc.)





EXPECTATION AND MOMENTS

Given a pdf f and a function g(X), the **expectation** $E_f(g(X))$ of g under f is the **weighted average**

$$E_f(g(X)) = \int_{\Omega} g(X)f(X) dX$$
, where $\Omega = \text{dom}(f)$.

The **moments** of a distribution are defined as

$$m_i = E(X^i)$$
, for $i = 0, ...,$

Note that $m_0 = 1$, by definition. The **mean** and **variance** of the distribution are given by $m_1 = E(X)$ and $m_2 - m_1^2 = E(X^2) - (E(X))^2$, respectively.

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distribution	$\mathbf{pdf} f(x)$	mean	variance	notes
uniform U(a, b)	$\frac{1}{b-a} \text{for } a \le x \le b$	$\frac{a+b}{2}$	$\frac{(b-a)^2}{12}$	most languages provide rand # generators for $U(a, b)$; used to generate r.v. with other distributions
Gaussian $N(\mu, \sigma^2)$	$\frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2} \text{ for } x \in \mathbb{R}$	μ	σ^2	if $X \sim N(\mu, \sigma^2)$, then $\frac{X-\mu}{\sigma} \sim N(0,1)$ (and vice- versa); very commonly used
Poisson $P(\lambda), \lambda \ge 0$	$\frac{\lambda^x}{x!}e^{-\lambda}$ for $x = 0,1,2,$	λ	λ	estimates the # of events that occur in a continuous time interval (# of calls received in 1-hour intervals)
binomial $\mathcal{B}(N, p), N \in \mathbb{N},$ $p \in [0,1]$	$\binom{N}{x} p^{x} (1-p)^{N-x}$ for $x = 0, 1,, N$	Np	Np(1-p)	describes the probability of exactly <i>x</i> successes in <i>N</i> independent trials if the probability of a success in a single trial is <i>p</i> (# of heads in <i>N</i> coin tosses)
$\frac{\log\text{-normal}}{\Lambda(\mu,\sigma^2)}$	$\frac{1}{x\sigma\sqrt{2\pi}}e^{-\frac{1}{2}\left(\frac{\ln x-\mu}{\sigma}\right)^2}$ for $x > 0$	$e^{(\mu+\sigma^2/2)}$	$e^{(2\mu+\sigma^2)} [e^{\sigma^2}-1]$	if $\ln X \sim N(\mu, \sigma^2)$, then $X \sim \Lambda(\mu, \sigma^2)$ (and <i>vice-versa</i>); positively skewed

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JOINT DISTRIBUTIONS

Univariate distributions are useful modeling tools, especially when the variables under consideration are **independent**.

In practice, that is not usually the case. A **joint distribution** $P(X_1, ..., X_n)$ gives the probability that each of $X_1, ..., X_n$ falls in a given range. The **multivariate normal distribution** $N(\mu, \Sigma)$ has pdf

$$f(x_1, \dots, x_n) := f(\boldsymbol{x}) = \frac{1}{\sqrt{(2\pi)^n \det(\boldsymbol{\Sigma})}} \exp\left\{-\frac{1}{2}(\boldsymbol{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1}(\boldsymbol{x} - \boldsymbol{\mu})\right\}$$

where μ is the mean vector and Σ the covariance matrix.

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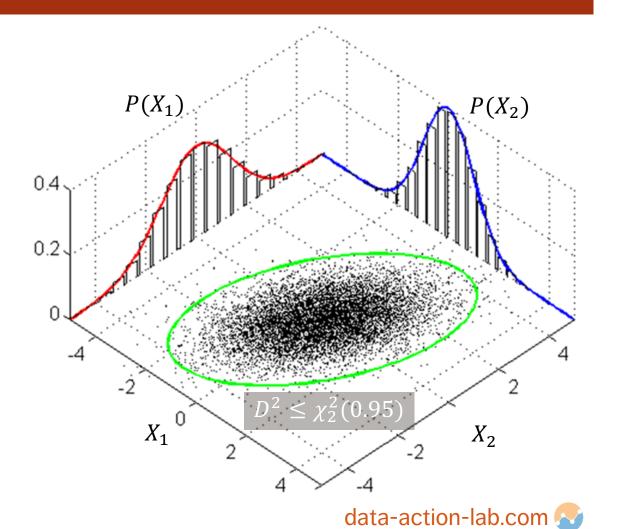


JOINT DISTRIBUTIONS

If Σ is positive definite, the multivariate normal is **non-degenerate**.

 $D = \sqrt{(x - \mu)^T \Sigma^{-1} (x - \mu)}$ is the **Mahalanobis** distance.

To generate a sample *x* from $N(\mu, \Sigma)$, let $z \sim N(0, I)$ and set $x = \mu + Az$, where $AA^T = \Sigma$ is the Cholesky decomposition.





Write R and/or Python code that lets you draw "random" samples from the various distributions discussed in this section.





CENTRAL LIMIT THEOREM

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LEARNING OBJECTIVES

What is the central limit theorem?

When is the central limit theorem relevant?





NORMAL DISTRIBUTION

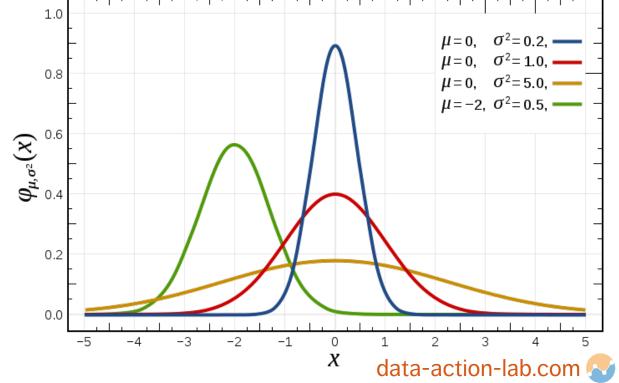
 $N(\mu, \sigma^2)$ is **fully characterized** by the mean μ and the standard deviation σ , which reduces estimation requirements.

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The probability of a value being drawn can be obtained if we know how many multiples of σ separate it from μ

- within σ from μ : $\approx 68\%$
- within 2σ from μ : $\approx 95\%$
- within 3σ from μ : $\approx 99.7\%$

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NORMAL DISTRIBUTION

The normal distribution is best suited for data meeting the following minimum requirements:

- strong tendency for the data to take on a central value
- positive, negative deviations from this central value are equally likely
- frequency of the deviations falls off rapidly as we move further away from the central value.

Symmetry of deviations leads to zero **skewness**; low prob. of large deviations from the central value leads to no **kurtosis**.

Its omnipresence in human affairs is linked to the **Central Limit Theorem**.



CENTRAL LIMIT THEOREM

Let $x_1, x_2, ..., x_n$ be a **random sample** from any (?) distribution with mean μ and variance σ^2 . If the sample observations are **independent** of each other, then the distribution of the average

$$w = \frac{x_1 + x_2 + \dots + x_n}{n}$$

is **approximately normal** (when $n \rightarrow \infty$) with mean and variance

$$\mu_w = \frac{1}{n}E(x_1 + \dots + x_n) = \mu, \ \sigma_w^2 = \frac{1}{n^2}E(x_1 + \dots + x_n - n\mu)^2 = \frac{1}{n}\sigma^2.$$

The CLT plays an important role in the prevalence of the normal distribution in human affairs.

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HOW LARGE IS LARGE?

If the underlying population is **normal**, the distribution of the sample mean is also **normal**, no matter the sample size n.

If the underlying population is **approximately symmetric**, the distribution of the sample mean is **approximately normal** for small sample sizes *n*.

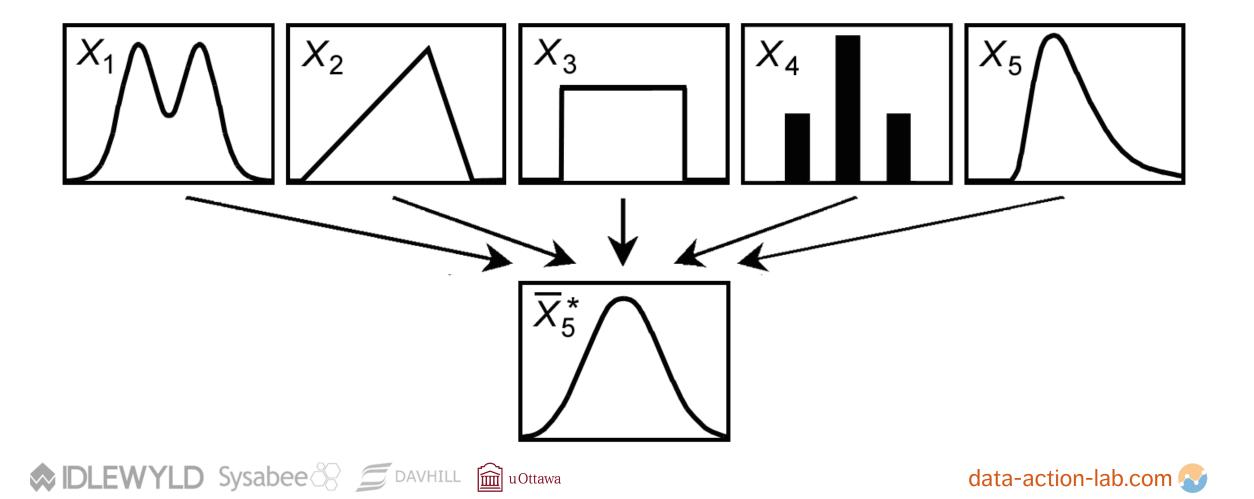
If the sample populations are **skewed** (or **disparate**), the sample size must typically reach 30 before the distribution of the sample mean becomes **approximately normal**.





[https://www.value-at-risk.net/central-limit-theorem]

CENTRAL LIMIT THEOREM IN ACTION



EXERCISES

A large freight elevator can transport a maximum of 9800 lbs. Suppose a load containing 49 boxes must be transported. From experience, the weight of boxes follows a distribution with mean $\mu = 205$ lbs and standard deviation $\sigma = 15$ lbs.

Using R and/or Python, estimate the probability that all 49 boxes can be safely loaded onto the freight elevator and transported.





ESTIMATION

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LEARNING OBJECTIVES

What is estimation, in a statistical sense?

What is estimation used for?

What is bias, in a statistical sense?



ESTIMATION

One of the goals of statistics is to try to **understand a large population** on the basis of the information available in a small sample.

In particular, we are interested in the population **parameters**, which are estimated using suitable sample statistics.

For example, we may use the **sample mean** $\bar{x} = \frac{1}{n} \sum_{i=1}^{n} x_i$ as an estimate for the true **population mean** μ .

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ESTIMATION

The **estimator** is a random variable; the **estimate** is a number.

As an another example, the **sample standard deviation** S is an estimator of the true **population standard deviation** σ and the computed value

$$s = \sqrt{\frac{1}{n-1}\sum_{i=1}^{n}(x_i - \bar{x})^2}$$

of *S* is an estimate of σ .

An estimator W of ω is **unbiased** if $E(W) = \omega$.

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BASIC MATHEMATICAL CONCEPTS

Let $X_1, ..., X_n$ be **random variables**, $b_1, ..., b_n \in \mathbb{R}$, and E, V, Cov be the **expectation**, **variance**, and **covariance** operators, respectively, i.e.:

- $E(X_i) = \mu_i$
- $\operatorname{Cov}(X_i, X_j) = \operatorname{E}(X_i X_j) \operatorname{E}(X_i) \operatorname{E}(X_j)$
- $V(X_i) = Cov(X_i, X_i) = E(X_i^2) E^2(X_i) = E(X_i^2) \mu_i^2 = \sigma_i^2$ and

$$E\left(\sum_{i=1}^{n} b_i X_i\right) = \sum_{i=1}^{n} b_i E(X_i) = \sum_{i=1}^{n} b_i \mu_i$$
$$V\left(\sum_{i=1}^{n} b_i X_i\right) = \sum_{i=1}^{n} b_i^2 V(X_i) + \sum_{i\neq j} b_i b_j \operatorname{Cov}(X_i, X_j)$$

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BASIC MATHEMATICAL CONCEPTS

The **bias** of an estimate is the average of the error in the estimate if the study is repeated many times independently under the same conditions.

The **variability** of an estimate is the extent to which the estimate would vary about its average value in the ideal scenario described above.

The **mean square error** of an estimate is a measure of the error that incorporates both elements:

$$MSE(\hat{\beta}) = V(\hat{\beta}) + Bias^2(\hat{\beta}),$$

where $\hat{\beta}$ is an estimator of β .

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BASIC MATHEMATICAL CONCEPTS

If the estimate $\hat{\beta}$ is unbiased, $E(\hat{\beta} - \beta) = 0$, then an approximate **95% confidence** interval (95% CI) for β is given approximately by

$$\hat{\beta} \pm 2\sqrt{\hat{V}(\hat{\beta})},$$

where $\widehat{V}(\hat{\beta})$ is a **sampling design-specific** estimate of $V(\hat{\beta})$.

But what is a 95% Cl, exactly?

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EXERCISE

The total time to manufacture a specific component is known to follow a normal distribution, for which the mean μ and variance σ^2 are not known. In an experiment, 10 components are manufactured; the sample time is given as following:

	1	2	3	4	5	6	7	8	9	10
Time	63.8	60.5	65.3	65.7	61.9	68.2	68.1	64.8	65.8	65.4

What are the best estimates for μ and σ^2 ? Provide a 95% CI for μ .

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LEARNING OBJECTIVES

What is a conditional probability and when is it useful?

What are some mathematical rules that govern probability?

What is Bayes' Theorem and when is it useful?





CONDITIONAL PROBABILITIES

We are often interested in the likelihood of an event occurring **given that another has occurred**.

Examples include:

- the probability that a train arrives on time given that it left on time
- the probability that a PC crashes given the operating system installed
- the probability that a bit is transmitted over a channel is received as a 1 given that the bit transmitted was a 1
- the probability that a website is visited given its number of in-links

Questions of this type are handled using conditional probability.

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CONDITIONAL PROBABILITIES

A **conditional probability** is the probability of an event taking place given that another event occurred.

The conditional probability of A given B, P(A|B), is defined as

$$P(A|B) = \frac{P(A \cap B)}{P(B)}$$

The probability that two events *A* and *B* both occur is obtained by applying the multiplication rule:

$$P(A \cap B) = P(B) P(A|B) = P(A) P(B|A)$$

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CONDITIONAL PROBABILITIES

Example (a classic): a family has two children (not twins). What is the probability that the youngest child is a girl given that at least one of the children is a girl? Assume that boys and girls are equally likely to be born.

Solution: Let *A* and *B* be the events that the youngest child is a girl and that at least one child is a girl, respectively:

$$A = \{GG, BG\}, \qquad B = \{GG, BG, GB\}$$

Then $P(A|B) = \frac{P(A \cap B)}{P(B)} = \frac{2}{3}$ (not $\frac{1}{2}$, as one might naively assume).

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RULES OF PROBABILITY

Let *I* denote relevant background information; *X*, *Y*, *Y*_k denote propositions, and -X denote the proposition that *X* is false.

The **plausibility** of X given I is denoted by P(X|I), ranging from 0 (false) to 1 (true).

Sum Rule: P(X|I) + P(-X|I) = 1Product Rule: $P(X, Y|I) = P(X|Y, I) \times P(Y|I)$ Bayes' Theorem: $P(X|Y, I) \times P(Y|I) = P(Y|X, I) \times P(X|I)$

Marginalization Rule: $P(X|I) = \sum P(X, Y_k|I)$, where $\{Y_k\}$ are exhaustive, disjoint

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The sum rule and the product rules are the **basic rules of probability**.

Bayes' Theorem and the **Marginalization Rule** are simple corollaries of these basic rules.

Bayes' Theorem is sometimes written is a slightly different form

$$P(X|Y,I) = \frac{P(Y|X,I) \times P(X|I)}{P(Y|I)}$$

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Set-up: assume that an experiment has been conducted to determine the degree of validity of a particular hypothesis, and that experimental data has been collected.

The central data analysis question: given everything that was known prior to the experiment, does the collected data support (or invalidate) the hypothesis?

Throughout, let *X* denote the proposition that the hypothesis in question is true, let *Y* denote the proposition that the experiment yielded the actual observed data, let *I* denote (as always) the relevant background information.

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Central data analysis question (reprise):

What is the value of *P*(hypothesis is true | observed data, *I*)?

Problem: this is nearly always impossible to compute directly.

Solution: using Bayes' Theorem,

 $P(\text{hypothesis} \mid \text{data}, I) = \frac{P(\text{data} \mid \text{hypothesis}, I) \times P(\text{hypothesis} \mid I)}{P(\text{data} \mid I)},$

it may be that the terms on the right are easier to compute.

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In the vernacular: the probability

- *P*(hypothesis | *I*) of the hypothesis being true prior to the experiment is the **prior**
- P(hypothesis | data, I) of the hypothesis being true once the experimental data is taken into account is the **posterior**
- P(data | hypothesis, I) of the experimental data being observed assuming that the hypothesis is true is the **likelihood**
- P(data | I) of the experimental data being observed independently of any hypothesis is the evidence

A given hypothesis includes a (potentially implicit) model which can be used to compute or approximate the **likelihood**.





Determining the **prior** is a source of considerable controversy

- conservative estimates (uninformative priors) often lead to reasonable results
- in the absence of information, go with maximum entropy prior

The **evidence** is harder to compute on theoretical grounds – evaluating the probability of observing data requires access to some model as part of *I*. Either

- that model was good, so there's no need for a new hypothesis
- that model was bad, so we dare not trust our computation





Thankfully, the evidence is rarely required on problems of parameter estimation (although it is crucial for model selection):

- prior to the experiment, there are numerous competing hypotheses
- the priors and likelihoods will differ, but not the evidence
- the evidence is not needed to differentiate the various hypotheses

Bayes' Theorem is often presented as

 $P(\text{hypothesis} | \text{data}, I) \propto P(\text{data} | \text{hypothesis}, I) \times P(\text{hypothesis} | I)$

or simply as posterior \propto likelihood×prior, that is to say, **beliefs should be updated in the presence of new information**.

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EXERCISE

Suppose that a test for a particular disease has a very high success rate. If a patient

- has the disease, the test accurately reports a 'positive' with probability 0.99;
- does not have the disease, the test accurately reports a 'negative' with probability 0.95.

Assume further that only 0.1% of the population has the disease. What is the probability that a patient who tests positive does not in fact have the disease?





MATRIX ALGEBRA

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Neo: What is the Matrix? **Trinity:** The answer is out there, Neo. It's looking for you, and it will find you if you want it to.

(The Matrix, the Wachowski Sisters)



LEARNING OBJECTIVES

What is the main mathematical object involved in linear algebra?

Why are matrices relevant in data science/data analysis?

What are some matrix operations?





LINEAR ALGEBRA

A **matrix** is an important mathematical tool that allows for easy organization of information, simplifies notation, and facilitates the application of algorithms to data.

Most statistical tools require **rectangular** data:

- each column contains a variable (feature, field, attribute)
 - indicator, target, question in a survey, etc.
- each row contains an **observation** (case, unit, item)
 - country, survey respondent, subject in an experiment, etc.
- each cell contains a **value** (measurement) for a particular variable and observation
 - GDP per capita for Canada, answer to a specific question, age, etc.





A matrix is a rectangular grid of **elements** arranged into **rows** and **columns**.

Matrices are often used in algebra to solve for unknown values in linear equations, and in geometry.

Matrix Addition: matrices can be added together (**element-wise**) as long as their **dimensions** are the same (i.e. both matrices have the same number of rows and columns), like so:

$$\begin{bmatrix} 3 & -2 \\ 4 & 1 \end{bmatrix} + \begin{bmatrix} 4 & 6 \\ 8 & 3 \end{bmatrix} = \begin{bmatrix} 7 & 4 \\ 12 & 4 \end{bmatrix}$$

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Multiplying a Matrix by a Scalar: a matrix of any dimension can be multiplied by a scalar by multiplying each element by the scalar

$$-1 \times \begin{bmatrix} 2 & 1 \\ 3 & -5 \\ 4 & 6 \end{bmatrix} = \begin{bmatrix} -2 & -1 \\ -3 & 5 \\ -4 & -6 \end{bmatrix}$$

Multiplying Matrices: two matrices *A* and *B* can be multiplied if their dimensions are **compatible** (i.e.,dim(A) = $n \times p$ and dim(B) = $p \times k$). The product C = AB is such that dim(C) = $n \times k$.

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The element in the *i*th row and *j*th column of the product C = AB is given by

$$c_{i,j} = a_{i,1}b_{1,j} + \dots + a_{i,p}b_{p,j}$$

For 2×2 matrices, this reduces to

$$\begin{bmatrix} a & b \\ c & d \end{bmatrix} \times \begin{bmatrix} e & f \\ g & h \end{bmatrix} = \begin{bmatrix} ae + bg & af + bh \\ ce + dg & cf + dh \end{bmatrix}$$

For instance,

$$A = \begin{bmatrix} 4 & 2 & 1 \\ 3 & 0 & 5 \end{bmatrix}, B = \begin{bmatrix} -2 \\ 3 \\ 0 \end{bmatrix} \Longrightarrow AB = \begin{bmatrix} 4 \times (-2) + 2 \times 3 + 1 \times 0 \\ 3 \times (-2) + 0 \times 3 + 5 \times 0 \end{bmatrix} = \begin{bmatrix} -2 \\ -6 \end{bmatrix}$$

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Transposing a Matrix: swapping the rows and the columns of a matrix is called **transposing** the matrix – it's denoted with a 'T':

$$\begin{bmatrix} 6 & 0 & -2 \\ 2 & 1 & 3 \end{bmatrix}^{\mathrm{T}} = \begin{bmatrix} 6 & 2 \\ 0 & 1 \\ -2 & 3 \end{bmatrix}$$

When applied to a data frame, transposing has the effect of interchanging the role of cases and observations.

For square matrices of size n (i.e. dim = $n \times n$), there are two special matrices: the **null matrix** 0_n (consisting only of zeroes), and the **identity matrix** I_n (diagonal entries are 1, all others 0).





For square matrices, two quantities often end up playing a fundamental role: the **trace** and the **determinant**.

The **trace** is the sum of the elements on the main diagonal:

$$\operatorname{tr} \begin{bmatrix} a_{11} & \cdots & a_{1n} \\ \vdots & \ddots & \vdots \\ a_{n1} & \cdots & a_{nn} \end{bmatrix} = a_{11} + a_{22} + \cdots + a_{nn}$$

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The **determinant** can be computed recursively. Let *A* be $n \times n$.

1. For
$$n = 1$$
, det $[a] = a$;

2. For
$$n = 2$$
, det $\begin{bmatrix} a_{11} & a_{12} \\ a_{21} & a_{22} \end{bmatrix} = a_{11}a_{22} - a_{12}a_{21}$

3. For a general *n*, let $D_{i,j}(A)$ be the determinant of the $(n-1) \times (n-1)$ matrix obtained by deleting the *i*th row and the *j*th column of *A*. The **Laplace expansion** of det *A* along the first column is

$$(-1)^{1+1}a_{11} D_{1,1}(A) + (-1)^{2+1}a_{21} D_{2,1}(A) + \dots + (-1)^{j+1}a_{j1} D_{j,1}(A) + \dots + (-1)^{n+1}a_{n1} D_{n,1}(A).$$

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The determinant can be expanded along any row/column without changing its value: $\begin{bmatrix} 1 & 0 \\ 2 \end{bmatrix}$

$$det \begin{bmatrix} 1 & 0 & -2 \\ 4 & -2 & 6 \\ 10 & 8 & 0 \end{bmatrix} = 1 \times det \begin{bmatrix} -2 & 6 \\ 8 & 0 \end{bmatrix} - 0 \times det \begin{bmatrix} 4 & 6 \\ 10 & 0 \end{bmatrix} + (-2) \times det \begin{bmatrix} 4 & -2 \\ 10 & 8 \end{bmatrix} = -152$$
or
$$det \begin{bmatrix} 1 & 0 & -2 \\ -2 & 6 \\ 10 & 8 & 0 \end{bmatrix} = -0 \times det \begin{bmatrix} 4 & 6 \\ 10 & 0 \end{bmatrix} + (-2) \times det \begin{bmatrix} 1 & -2 \\ 10 & 0 \end{bmatrix} - 8 \times det \begin{bmatrix} 1 & -2 \\ 4 & 6 \end{bmatrix} = -152$$
and
$$tr \begin{bmatrix} 1 & 0 & -2 \\ 4 & -2 & 6 \\ 10 & 8 & 0 \end{bmatrix} = 1 + (-2) + 0$$

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The determinant is linked to the **inverse** of a matrix.

In number arithmetic every number $a \neq 0$ has an inverse *b* written as a^{-1} or 1/a such that ba = ab = 1. Similarly a square matrix *A* may have an inverse $B = A^{-1}$ where $AB = BA = I_n$.

Miscellaneous:

- Non-square matrices do not possess inverses.
- Not all square matrices have an inverse (only those with $det(A) \neq 0$).
- A matrix which has an inverse is said to be **non-singular**.

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If
$$ad - bc \neq 0$$
 then the matrix $A = \begin{bmatrix} a & b \\ c & d \end{bmatrix}$ has a (unique) inverse:
$$A^{-1} = \frac{1}{ad - bc} \begin{bmatrix} d & -b \\ -c & a \end{bmatrix}$$

For n > 2, other computation methods exist, such as **Gaussian elimination:** if a sequence of row operations $(yR_j + xR_i \rightarrow R_j, R_j \leftrightarrow R_i)$ applied to a square matrix A reduce it to the identity matrix I of the same size, then the same sequence of operations applied to I reduces it to A^{-1} .

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If we cannot reduce A to I then A^{-1} does not exist. This will become evident by the appearance of a row of zeros. There is no unique route from A to I and it is experience which selects the optimal route.

It is more efficient to do the two reductions simultaneously;

$$\begin{bmatrix} A|I \end{bmatrix} = \begin{bmatrix} 1 & 3 & 3|1 & 0 & 0 \\ 1 & 4 & 3|0 & 1 & 0 \\ 2 & 7 & 7|0 & 0 & 1 \end{bmatrix} \xrightarrow{R_2 - R_1 \to R_2} \begin{bmatrix} 1 & 3 & 3|1 & 0 & 0 \\ 0 & 1 & 0| -1 & 1 & 0 \\ 0 & 1 & 1| -2 & 0 & 1 \end{bmatrix}$$

$$\xrightarrow{R_1 - 3R_2 \to R_1} \begin{bmatrix} 1 & 0 & 3| & 4 & -3 & 0 \\ 0 & 1 & 0| -1 & 1 & 0 \\ 0 & 0 & 1| -1 & -1 & 1 \end{bmatrix} \xrightarrow{R_1 - 3R_3 \to R_1} \begin{bmatrix} 1 & 0 & 0| & 7 & 0 & -3 \\ 0 & 1 & 0| -1 & 1 & 0 \\ 0 & 0 & 1| -1 & -1 & 1 \end{bmatrix} = \begin{bmatrix} I|A^{-1} \end{bmatrix}$$

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EXERCISES

In R, construct 3×3 square matrices A, B, C and compute the following:

- $A + B, BC, CB, A^{T}, CA^{T}$
- tr(A), tr(3A), tr(C), tr(-C), tr(3A C)
- $det(A), det(A^T), det(B), det(C), det(BC)$
- A^{-1}, B^{-1}, C^{-1} , if the respective determinants are $\neq 0$
- $det(A^{-1})$, $det(B^{-1})$, $det(C^{-1})$, if the respective matrices are invertible

Can you infer rules from these computations?

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EIGENVALUES AND EIGENVECTORS

STATISTICAL AND MATHEMATICAL FOUNDATIONS





LEARNING OBJECTIVES

What is an eigenvalue?

What is an eigenvector?

What is a use case for these mathematical concepts?





EIGENVECTORS AND EIGENVALUES

An **eigenvector** of a matrix A is a vector $v \neq 0$ such that, for some scalar λ , $Av = \lambda v$.

The value λ is called an **eigenvalue** of *A* associated with ν .

The eigenvalues of an $n \times n$ matrix A satisfy $det(A - \lambda I_n) = 0$. The left-hand side is a polynomial in λ , and is called the **characteristic polynomial** of A, denoted by $p_A(\lambda)$.

To find the eigenvalues of *A*, we find the roots of $p_A(\lambda)$.

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EXAMPLE

Let $A = \begin{bmatrix} 2 & -4 \\ -1 & -1 \end{bmatrix}$. Then $p_A(\lambda) = \det(A - \lambda I) = (\lambda - 3)(\lambda + 2)$. Thus, $\lambda_1 = 3$ and $\lambda_2 = -2$ are the eigenvalues of A.

To find eigenvectors corresponding to an eigenvalue λ , we solve the system of linear equations given by $(A - \lambda I)v = 0$.

Let's find the eigenvectors corresponding to $\lambda_1 = 3$, by solving

$$(A-3I)\boldsymbol{\nu} = \begin{bmatrix} 2-3 & -4 \\ -1 & -1-3 \end{bmatrix} \begin{bmatrix} v_1 \\ v_2 \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix}.$$

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EXAMPLE

This yields the following equations:

$$-v_1 - 4v_2 = 0 , \qquad -v_1 - 4v_2 = 0$$

If we let $v_2 = t$, then $v_1 = -4t$, and so all eigenvectors corresponding to $\lambda_1 = 3$ are multiples of $\begin{bmatrix} -4\\1 \end{bmatrix}$.

A similar computation shows that all eigenvectors corresponding to $\lambda_2 = -2$ are multiples of $\begin{bmatrix} 1 \\ 1 \end{bmatrix}$.

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EIGEN-DECOMPOSITION

If an $n \times n$ matrix A has n linearly independent eigenvectors, then A may be **decomposed** in the following manner:

$$A = BAB^{-1},$$

where Λ is a diagonal matrix whose diagonal entries are the eigenvalues of A and the columns of B are the corresponding eigenvectors of A.





EXAMPLE

We have seen that the eigenvalues of $A = \begin{bmatrix} 2 & -4 \\ -1 & -1 \end{bmatrix}$ are $\lambda_1 = 3$ and $\lambda_2 = -2$, and that the corresponding eigenvectors are $\begin{bmatrix} -4 \\ 1 \end{bmatrix}$ and $\begin{bmatrix} 1 \\ 1 \end{bmatrix}$.

Thus,
$$A = \begin{bmatrix} 3 & 0 \\ 0 & -2 \end{bmatrix}$$
, $B = \begin{bmatrix} -4 & 1 \\ 1 & 1 \end{bmatrix}$, and

$$A = \begin{bmatrix} -4 & 1 \\ 1 & 1 \end{bmatrix} \begin{bmatrix} 3 & 0 \\ 0 & -2 \end{bmatrix} \begin{bmatrix} -4 & 1 \\ 1 & 1 \end{bmatrix}^{-1} = \begin{bmatrix} -4 & 1 \\ 1 & 1 \end{bmatrix} \begin{bmatrix} 3 & 0 \\ 0 & -2 \end{bmatrix} \frac{1}{-4 \times 1 - 1 \times 1} \begin{bmatrix} 1 & -1 \\ -1 & -4 \end{bmatrix}$$

$$= \begin{bmatrix} -4 & 1 \\ 1 & 1 \end{bmatrix} \begin{bmatrix} 3 & 0 \\ 0 & -2 \end{bmatrix} \begin{bmatrix} -1/5 & 1/5 \\ 1/5 & 4/5 \end{bmatrix}$$
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Compute the eigen-decomposition of the matrices *A*, *B*, *C* you constructed in the previous module.





OPTIMIZATION

STATISTICAL AND MATHEMATICAL FOUNDATIONS





LEARNING OBJECTIVES

What is optimization?

When is optimization useful?

What is a cost function?

Why are minima and maxima relevant to optimization?

What are techniques that can be used to carry out optimization?





OPTIMIZATION

Suppose we have a **cost** (objective) function $f: \mathbb{R}^n \to \mathbb{R}$ to **optimize** (the maximum likelihood function of linear regression, for instance).

Seeking a maximum for f is equivalent to seeking a minimum for -f.

The aim is to find parameter values *x* that minimize this function:

$$\boldsymbol{x}^* = \arg\min_{\boldsymbol{x}} f(\boldsymbol{x})$$

The cost function could be subjected to a number of constraints

$$c_i(\mathbf{x}) = 0, i = 1, \dots, m; c_j(\mathbf{x}) \ge 0, j = 1, \dots, k; \mathbf{x} \in \Omega \subseteq \mathbb{R}^n.$$

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OPTIMIZATION

The optimization problem can be viewed as a **decision problem** that involves finding the "best" vector **x** over all possible vectors in $\Omega \subseteq \mathbb{R}^n$.

This vector is called the **minimizer** of f over Ω . There may be multiple minimizers, or none.

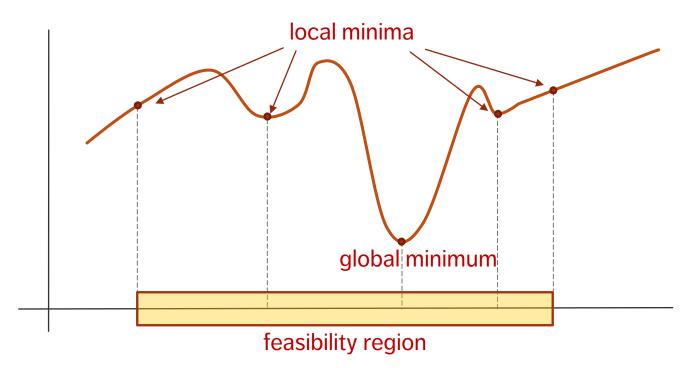
If $\Omega = \mathbb{R}^n$, then we refer to the problem as an **unconstrained** optimization problem.

In general, this is not a trivial problem (consult the literature).





TYPE OF MINIMA



In many instances, optimization is a **numerical** endeavour. Which of the minima is found depends on the algorithm's **starting point**.



GOLDEN SECTION METHOD

The **golden section search** is a technique for finding the extremum (minimum or maximum) of a strictly unimodal function by successively narrowing the range of values inside which the extremum is known to exist.

The technique derives its name from the fact that the algorithm maintains the function values for triples of points whose distances form a **golden ratio**.





GOLDEN SECTION METHOD

Let [a, b] be the interval of the current bracket (i.e. the optimizer resides in [a, b]), and assume f(a), f(b) have already have been computed. Denote $\varphi = (1 + \sqrt{5})/2$.

1. Let
$$c = b - \frac{(b-a)}{\varphi}$$
, $d = a + \frac{(b-a)}{\varphi}$;

- 2. If f(c), f(d) are not available, compute them;
- 3. If f(c) < f(d) (to find a min to find a max, reverse the order) then move the data: $(b, f(b)) \leftarrow (d, f(d))$ and $(d, f(d)) \leftarrow (c, f(c))$ and update $c = b - (a - b)/\varphi$ and f(c);
- 4. Otherwise, move the data $(a, f(a)) \leftarrow (c, f(c))$ and $(c, f(c)) \leftarrow (d, f(d))$ and update $d = a + (b a)/\varphi$ and f(d);

5. The interval [*c*, *d*] brackets the optimizer. Continue until tolerance is reached.

NEWTON'S METHOD

In calculus, we learn that a function $f: \Omega \subseteq \mathbb{R}^n \to \mathbb{R}$ which is sufficiently well behaved reaches its max/min either at a **critical point** (i.e. where $\nabla f = \mathbf{0}$) or on the **domain boundary** $\partial \Omega$.

Thus, to identify candidate optimizers, we must be able to solve general systems of the form g(x) = 0.

Newton's Method is a powerful method for finding roots of functions.





NEWTON'S METHOD

For n = 1, the algorithm is shown below (it is quite similar in the general case).

Let x = c be an (unknown) zero of a differentiable function f in an open interval containing c.

- **1.** make an initial approximation x_1 "close" to c
- 2. determine a new approximation using the formula $x_2 = x_1 \frac{f(x_1)}{f'(x_1)}$.
- 3. If $|x_2 x_1|$ is less than the desired accuracy (which needs to be specified), x_2 serves as the final approximation. Otherwise, return to step 2. and calculate a new approximation.

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Use the golden section method and Newton's method to find a root of

$$f(x) = e^{-x} \sin(x)$$
 and $g(x) = x \ln(x)$.

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SAMPLING THEORY AND STUDY DESIGN

DATA COLLECTION AND DATA PROCESSING

"The latest survey shows that 3 out of 4 people make up 75% of the population"

D. Letterman





THE GOAL OF GOOD STUDY/SAMPLING DESIGN

We need data that can:

- provide legitimate insight into our system of interest;
- provide correct, accurate answers to relevant questions;
- support the drawing of legitimate, valid conclusions, with the ability to qualify these conclusions in terms of scope and precision.

This starts with **study design** – what data to collect and how it should be collected





 "A Dartmouth graduate student used an MRI machine to study the brain activity of a salmon as it was shown photographs and asked questions. The most interesting thing about the study was not that a salmon was studied, but that the salmon was dead. Yep, a dead salmon purchased at a local market was put into the MRI machine, and some patterns were discovered. There were inevitably patterns—and they were invariably meaningless."





NPS AND PATTERN FISHING

Two separate issues can be combined to cause **problems** with data analysis:

- drawing conclusions (inferences) from a sample about a population that are not warranted by the sample collection method (symptomatic of NPS);
- looking for any available patterns in the data and then coming up with post hoc explanations for these patterns.

Alone or in combination, these lead to poor (and **potentially harmful**) conclusions.





STUDIES AND SURVEYS

A **survey** is any activity that collects information about characteristics of interest:

- in an **organized** and **methodical** manner;
- from some or all **units** of a population;
- using well-defined concepts, methods, and procedures, and
- compiles such information into a meaningful summary form.





SAMPLING MODELS

A **census** is a survey where information is collected from all units of a population, whereas a **sample survey** uses only a fraction of the units.

When survey sampling is done properly, we may be able to use various **statistical methods** to make **inferences** about the **target population** by sampling a (comparatively) small number of units in the **study population**.





DECIDING FACTORS

In some instances, information about the **entire** population is required in order to answer questions, whereas in others it is not necessary. The **survey type** depends on multiple factors:

- the type of question that needs to be answered;
- the required precision;
- the cost of surveying a unit;
- the time required to survey a unit;
- size of the population under investigation, and
- the prevalence of the attributes of interest.

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STUDY/SURVEY STEPS

Studies or surveys follow the same general steps:

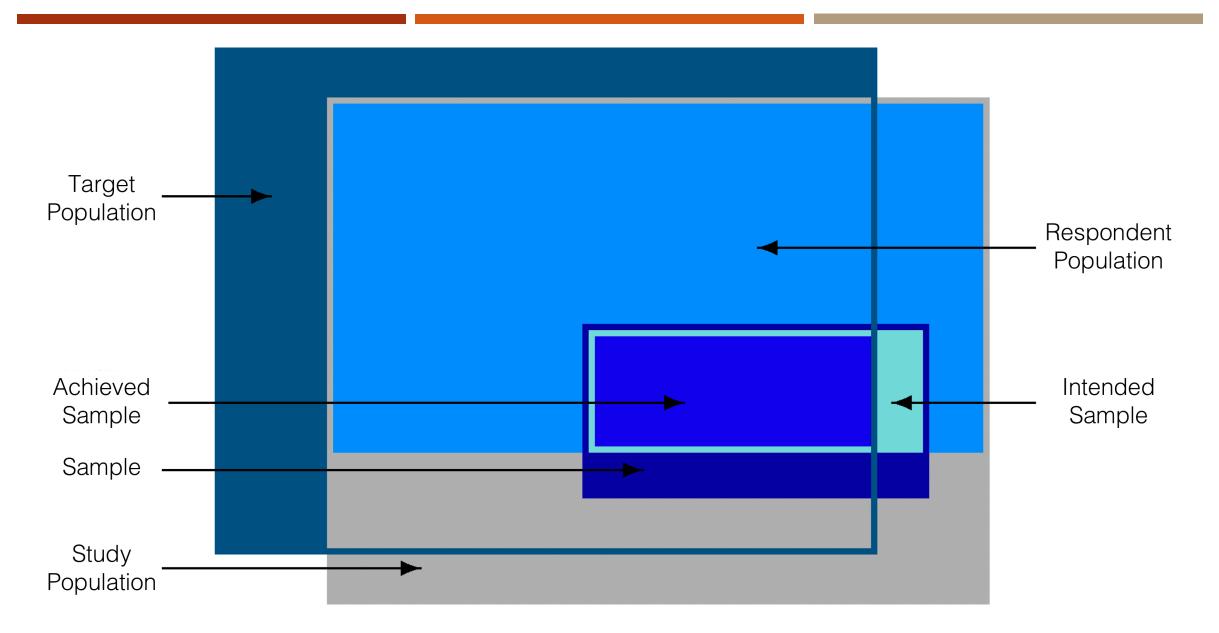
- 1. statement of objective
- 2. selection of survey frame
- 3. sampling design
- 4. questionnaire design
- 5. data collection
- 6. data capture and coding

- 7. data processing and imputation
- 8. estimation
- 9. data analysis
- 10. dissemination
- 11. documentation

The process is not always linear, but there is a definite movement from objective to dissemination.







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SURVEY FRAMES

The **frame** provides the means of **identifying** and **contacting** the units of the study population. It is generally costly to create and to maintain (in fact, there are organisations and companies that specialise in building and/or selling such frames).

Useful frames contain:

- identification data,
- contact data,
- classification data,
- maintenance data, and
- linkage data.





SURVEY FRAMES

The ideal frame contains identification data, contact data, classification data, maintenance data, and linkage data, and must minimize the risk of **undercoverage** or **overcoverage**, as well as the number of duplications and misclassifications (although some issues that arise can be fixed at the data processing stage).

A statistical sampling approach is contraindicated unless the selected frame is

- relevant (which is to say, it corresponds, and permits accessibility to, the target population),
- **accurate** (the information it contains is valid),
- timely (it is up-to-date), and
- competitively priced.





MODES OF DATA COLLECTION

Paper-based vs. computer-assisted

- self-administered questionnaires are used when the survey requires detailed information to allow the units to consult personal records; associated with high non-response rate.
- interviewer-assisted questionnaires use well-trained interviewers to increase the response rate and overall quality of the data; face-to-face vs. telephone.
- **computer-assisted interviews** combine data collection and data capture, which saves time.
- unobtrusive direct observation
- diaries to be filled (paper or electronic)
- omnibus surveys
- email, Internet, and social media





SURVEY ERROR



Statistical sampling can help provide estimates, but importantly, it can also provide some control over the **total error** (TE) of the estimates.

Ideally, TE = 0. In practice, there are two main contributions to TE: **sampling errors** (due to the choice of sampling scheme), and **nonsampling errors** (everything else).

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NONSAMPLING ERROR

Nonsampling error can be controlled, to some extent:

- coverage error can be minimized by selecting high quality, up-to-date survey frames;
- non-response error can be minimized by careful choice of the data collection mode and questionnaire design, and by using "call-backs" and "follow-ups";
- measurement error can be minimized by careful questionnaire design, pre-testing of the measurement apparatus, and cross-validation of answers.

In practice, these suggestions are not that useful in modern times (landline-based survey frames are becoming irrelevant due to demographics, response rates for surveys that are not mandated by law are low, etc.). This explains, in part, the over-use of **web scraping** and **non-probabilistic sampling**.





NONPROBABILISTIC SAMPLING

Nonprobabilistic sampling (NPS) methods (designs) select sampling units from the target population using subjective, non-random approaches.

- NPS are quick, relatively inexpensive and convenient (no survey frame required).
- NPS methods are ideal for exploratory analysis and survey development.

Unfortunately, NPS are often used instead of probabilistic designs (problematic)

- the associated selection bias makes NPS methods unsound when it comes to inferences (they
 cannot be used to provide reliable estimates of the sampling error, the only component of TE under
 the analyst's direct control);
- automated data collection often fall squarely in the NPS camp we can still analyze data collected with a NPS approach, but may not generalize the results to the target population.





NPS METHODS

Haphazard

man on the street, depends on availability of units and interviewer bias

Volunteer

self-selection bias

Judgement

biased by inaccurate preconceptions about the target population

Quota

exit polling, ignores non-response bias

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NPS METHODS

Modified

starts probabilistic, switches to quota as a reaction to high non-response rates

Snowball

• "pyramid" scheme

There are contexts where NPS methods might fit a client's or an organization's need (and that remains their decision to make, ultimately), but they must be informed of the drawbacks, and presented with some probabilistic alternatives.





PROBABILISTIC SAMPLING

Probabilistic sample designs are usually more **difficult** and **expensive** to set-up (due to the need for a quality survey frame), and take longer to complete.

They provide **reliable estimates** for the attribute of interest and the **sampling error**, paving the way for small samples being used to draw inferences about larger target populations (in theory, at least; the non-sampling error components can still affect results and generalisation).





BASIC MATHEMATICAL CONCEPTS

Consider a finite population \mathcal{U} , with N units and measurements $\{u_1, \dots, u_N\}$.

The mean and variance of the population for the variable of interest are given by

$$\mu = \frac{1}{N} \sum_{j=1}^{N} u_j, \qquad \sigma^2 = \frac{1}{N} \sum_{j=1}^{N} (u_j - \mu)^2,$$

If $\mathcal{Y} \subseteq \mathcal{U}$ is a **sample** of the population with *n* units and measurements $\{y_1, \dots, y_n\}$, then the **sample mean** and **sample variance** are given by

$$\bar{y} = \frac{1}{n} \sum_{i=1}^{n} y_i$$
, $s^2 = \frac{1}{n-1} \sum_{i=1}^{n} (y_i - \bar{y})^2$.

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BASIC MATHEMATICAL CONCEPTS

Let $X_1, ..., X_n$ be **random variables**, $b_1, ..., b_n \in \mathbb{R}$, and E, V, Cov be the **expectation**, **variance**, and **covariance** operators, respectively, i.e.:

- $E(X_i) = \mu_i$
- $\operatorname{Cov}(X_i, X_j) = \operatorname{E}(X_i X_j) \operatorname{E}(X_i) \operatorname{E}(X_j)$
- $V(X_i) = Cov(X_i, X_i) = E(X_i^2) E^2(X_i) = E(X_i^2) \mu_i^2 = \sigma_i^2$ and

$$E\left(\sum_{i=1}^{n} b_i X_i\right) = \sum_{i=1}^{n} b_i E(X_i) = \sum_{i=1}^{n} b_i \mu_i$$
$$V\left(\sum_{i=1}^{n} b_i X_i\right) = \sum_{i=1}^{n} b_i^2 V(X_i) + \sum_{i\neq j} b_i b_j \operatorname{Cov}(X_i, X_j)$$

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BASIC MATHEMATICAL CONCEPTS

The **bias** in an error component is the average of that error component if the survey is repeated many times independently under the same conditions. An **unbiased** estimate is one for which the bias is nil.

The **variability** in an error component is the extent to which that component would vary about its average value in the ideal scenario described above.

The **mean square error** of an error component is a measure of its size:

$$MSE(\hat{\beta}) = V(\hat{\beta}) + Bias^2(\hat{\beta}),$$

Where $\hat{\beta}$ is an estimator of β .

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CONFIDENCE INTERVALS

If the estimate $\hat{\beta}$ is unbiased, $E(\hat{\beta} - \beta) = 0$, then an approximate **95% confidence** interval (95% CI) for β is given approximately by

$$\hat{\beta} \pm 2\sqrt{\hat{\mathbb{V}}(\hat{\beta})}$$
,

where $\widehat{V}(\hat{\beta})$ is a **sampling design-specific** estimate of $V(\hat{\beta})$.

But what is a 95% Cl, exactly?



SAMPLING DESIGN

Different **sampling designs** have distinct advantages and disadvantages.

They can be used to compute estimates

- for various population attributes: mean, total, proportion, ratio, difference, etc.
- for the corresponding 95% Cl.

We might also want to compute sample sizes for a given **error bound** (an upper limit on the radius of the desired 95% CI), and how to determine the **sample allocation** (how many units to be sampled in various sub-population groups).

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SAMPLING DESIGN – UNIVERSE OF DISCOURSE

Target population:

N units and measurements $\mathcal{U} = \{u_1, \dots, u_N\}$

True population attributes:

• mean μ , variance σ^2 , total τ , proportion p

Sample population:

• *n* units and measurements $\mathcal{Y} = \{y_1, \dots, y_n\} \subseteq \mathcal{U}$

Sample population attributes:

sample mean \overline{y} , sample variance s^2 , sample total $\hat{\tau}$, sample proportion \hat{p}



PROBABILISTIC SAMPLING DESIGNS

Simple random sampling (SRS) Stratified random sampling (StS) Systematic sampling (SyS) Cluster sampling (CIS)

Probability proportional-to-size sampling (PPS) Replicated sampling (ReS) Multi-stage sampling (MSS) Multi-phase sampling (MPS)





SAMPLING DESIGN – UNIVERSE OF DISCOURSE

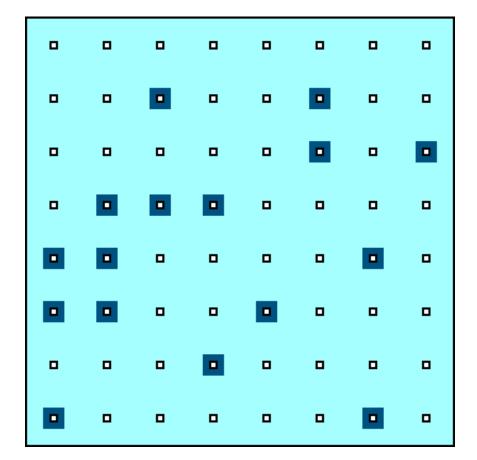
- **Goal:** estimate the true population attributes μ , σ^2 , τ , p via the sample population attributes \overline{y} , s^2 , $\hat{\tau}$, \hat{p} , n, and the size N of the target population.
- For a given characteristic, we define δ_i as 1 or 0 depending on whether the sample unit y_i possesses the characteristic in question or not.
- We use the error bound $B = 2\sqrt{\widehat{V}}$.

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SIMPLE RANDOM SAMPLING (SRS)

- In SRS, *n* units are selected randomly from the frame.
- Advantages:
 - easiest sampling design to implement
 - sampling errors are well-known and easy to estimate
 - does not require auxiliary information
- Disadvantages:
 - makes no use of auxiliary information
 - no guarantee that the sample is representative
 - costly if sample is widely spread out, geographically



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SRS ESTIMATORS

Estimators:

$$\overline{y} = \frac{1}{n} \sum_{i=1}^{n} y_i, \qquad \hat{\tau} = N \overline{y}, \qquad \hat{p} = \frac{1}{n} \sum_{i=1}^{n} \delta_i$$

Sample Design-Specific Variance Estimates:

$$\widehat{\mathbb{V}}(\overline{y}) = \frac{s^2}{n} \left(1 - \frac{n}{N}\right), \qquad \widehat{\mathbb{V}}(\widehat{\tau}) = N^2 \widehat{\mathbb{V}}(\overline{y}), \qquad \widehat{\mathbb{V}}(\widehat{p}) = \frac{\widehat{p}(1 - \widehat{p})}{n} \left(1 - \frac{n}{N}\right)$$

Sample Allocation:

$$n_{\bar{y}} = \frac{4N\tilde{\sigma}^2}{(N-1)B^2 + 4\tilde{\sigma}^2}, \qquad n_{\hat{\tau}} = \frac{4N^3\tilde{\sigma}^2}{(N-1)B^2 + 4N^2\tilde{\sigma}^2}, \qquad n_{\hat{p}} = \frac{4\tilde{p}(1-\tilde{p})}{(N-1)B^2 + 4\tilde{p}(1-\tilde{p})}$$

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STRATIFIED RANDOM SAMPLING (STS)

• In StS, $n = n_1 + \dots + n_k$ units are selected randomly from k frame **strata**.

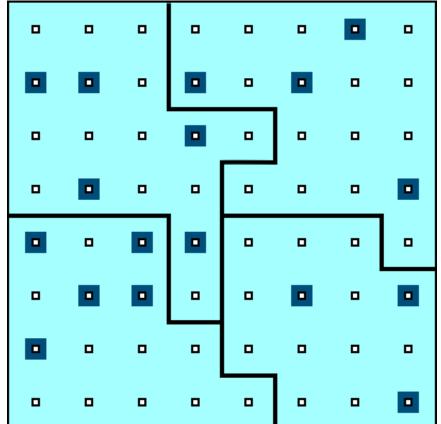
Advantages:

- may produce smaller error bound on estimation than SRS
- may be less expensive if elements are conveniently strat.
- may provide estimates for sub-populations

Disadvantages:

- no major disadvantage
- if there are no natural ways to stratify the frame into homogeneous groupings, StS is roughly equivalent to SRS

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STS ESTIMATORS

Estimators:

$$\bar{y}_{st} = \sum_{j=1}^{k} \frac{N_j}{N} \bar{y}_j, \quad \hat{\tau}_{st} = N \bar{y}_{st}, \qquad \hat{p}_{st} = \sum_{j=1}^{k} \frac{N_j}{N} \hat{p}_j$$

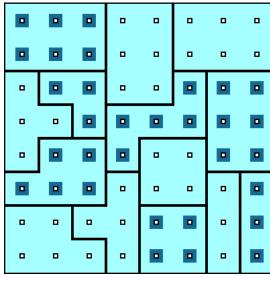
Sample Design-Specific Variance Estimates:

$$\widehat{\mathbb{V}}(\bar{y}_{st}) = \frac{1}{N^2} \sum_{j=1}^k N_j^2 \widehat{\mathbb{V}}(\bar{y}_j), \qquad \widehat{\mathbb{V}}(\hat{\tau}_{st}) = N^2 \widehat{\mathbb{V}}(\bar{y}_{st}), \qquad \widehat{\mathbb{V}}(\hat{p}_{st}) = \frac{1}{N^2} \sum_{j=1}^k N_j^2 \widehat{\mathbb{V}}(\hat{p}_j)$$

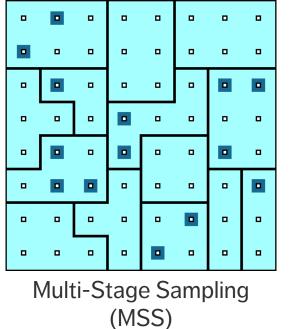
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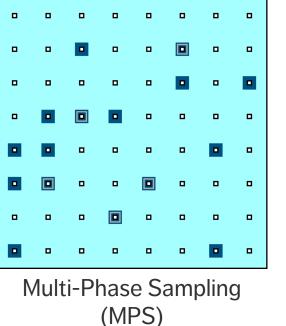
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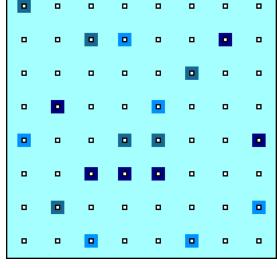
OTHER EXAMPLES OF SAMPLING DESIGNS



Cluster Sampling (CIS)







Replicated Sampling (ReS)

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EXERCISES

You are charged with estimating the yearly salary of data scientists in Canada.

Identify potential:

- populations (target, study, respondent, sampling frames)
- samples (intended, achieved)
- unit information (unit, response variate, population attribute)
- sources of bias (coverage, nonresponse, sampling, measurement) and variability (sampling, measurement).





CONFIDENCE INTERVALS (COMING SOON)





HYPOTHESIS TESTING (COMING SOON)





REGRESSION





LEARNING OBJECTIVES

What is regression modelling?

What are some types of regression modeling?

When is regression modeling useful?





REGRESSION MODELING

The most common data modeling methods are regressions, both linear and logistic

 ~90% of real data applications end up using a simple regression as their final model, typically after very careful data preparation, encoding, and creation of variables.

There are several reasons for their frequent use:

- generally straightforward to understand and to train
- mean square error (MSE) objective function has a closed-form linear solution
- system of equations can usually be solved through matrix inversion or linear manipulation





REGRESSION MODELING

The data structure of a general modeling task	X_1	X_2	 X_p	Y
is represented by	x_{11}	x_{12}	 x_{1p}	y_1
	x_{21}	x_{22}	 x_{2p}	y_2
We consider p independent variables X_i			 	
to try to predict the dependent variable Y .	x_{n1}	x_{n2}	 x_{np}	y_n

In order to simplify the discussion in the following, we introduce the matrix notation $X[n \times p], Y[n \times 1], \beta[p \times 1]$, where *n* is the # of observations and *p* is the # of independent variables.

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LINEAR REGRESSION

The basic assumption of linear regression is that the dependent variable y can be **approximated** by a linear combination of the independent variables as follows:

$$Y=X\beta+\varepsilon,$$

where $\beta \in \mathbb{R}^p$ is to be determined based on the training set, and for which

$$\mathsf{E}(\boldsymbol{\varepsilon}|\boldsymbol{X}) = 0, \qquad \mathsf{E}(\boldsymbol{\varepsilon}\boldsymbol{\varepsilon}^T|\boldsymbol{X}) = \sigma^2 I.$$

Typically, the errors are also assumed to be normally distributed, that is :

$$\boldsymbol{\varepsilon}|\boldsymbol{X} \sim N(0, \sigma^2 I).$$

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LINEAR REGRESSION

If $\hat{\beta}_i$ is the estimate of the true coefficient β_i , the **linear regression** model associated with the data is

$$\widehat{Y}(\mathbf{x}) = \widehat{\beta}_0 + \widehat{\beta}_1 x_1 + \dots + \widehat{\beta}_p x_p$$

In matrix form, the regression problem requires a solution $\hat{\beta}$ to the **normal equation** $X^T X \beta = X^T Y$.

When the symmetric positive definite matrix $X^T X$ is invertible, the fitted coefficient is simply $\hat{\beta} = (X^T X)^{-1} (X^T Y)$. Note that $X^T X$ is a $p \times p$ matrix, which makes the inversion "easier" to compute, relatively speaking, when n is large.

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GENERALIZED LINEAR REGRESSION

Generalized linear models (GLMs) extend linear statistical models by accommodating response variables with **non-normal** conditional distributions.

Except for the **error structure**, a GLM is essentially the same as for a linear model:

 $Y_i \sim \text{some distribution with mean } \mu_i$, where $g(\mu_i) = x_i^T \beta$

A GLM therefore consists of three parts:

- a **systematic** component $x_i^T \beta$
- a random component specified distribution for Y_i
- a **link** function *g*





GENERALIZED LINEAR REGRESSION

We could specify **any** distribution for the outcome variable *Y*...

 but the mathematics of GLM work nicely only for the exponential family of distributions (most common statistical distributions fall into this family: such as the normal, binomial, Poisson, gamma, and others).

Linear regression is an example of GLM:

- systematic component: $x_i^T \beta$
- random component: $Y_i \sim N(\mu_i, \sigma^2)$
- link: $g(\mu) = \mu$, the identity link

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EXAMPLE

In the early stages of an epidemic, the rate at which new cases occur increases exponentially through time.

If μ_i is the expected number of new cases on day t_i , a model taking the form

 $\mu_i = \gamma \exp(\delta t_i)$

might be appropriate. If we take the log of both sides, we get

$$log(\mu_i) = log(\gamma) + \delta t_i = \beta_0 + \beta_1 t_i = (1, t_i)^T (\beta_0, \beta_1).$$

systematic component

Furthermore, since the we measure the number of new cases (a count), the **Poisson** distribution could be a reasonable choice.

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ADVANTAGES OF GLM

No need to transform *Y* to have a normal distribution

Choice of link is **separate** from the choice of random component

more modeling flexibility

If link produces **additive effects**, no need for constant variance

Models are fitted via ML estimation

optimal properties of the estimators

Inference tools and model checks apply to other GLMs

Wald ratio test, likelihood ratio test, deviance, residuals, confidence intervals, etc.

See PROC GENMOD in SAS, or glm() in R

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EXERCISE

An auto part is manufactured by a company once a month in lots that vary in size as demand fluctuates. The data below represent observations on lot size (y), and number of employee-hours of labor (x) for 10 recent production runs.

Fit a simple regression model $y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$, where $E(\varepsilon_i) = 0$, $E(\varepsilon_i \varepsilon_j) = 0$ for $i \neq j$, and $V(\varepsilon_i) = \sigma^2$ if the observations are:

 $\boldsymbol{Y} = [73, 50, 128, 170, 87, 108, 135, 69, 148, 132]^{\mathrm{T}},$

 $\boldsymbol{x} = [30, 20, 60, 80, 40, 50, 60, 30, 70, 60]^{\mathrm{T}}.$

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